An Optimal Multiline TRL Calibration Algorithm

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Abstract — We examine the performance of two on-wafer multiline Thru-Reflect-Line (TRL) calibration algorithms: the popular multiline TRL algorithm implemented in the MultiCal® software package, and a newly implemented iterative algorithm designed to give optimal results in the presence of measurement noise. We show that the iterative algorithm outperforms the MultiCal software in the presence of measurement noise, and verify its uncertainty estimates.

I. INTRODUCTION

We compare the multiline Thru-Reflect-Line (TRL) vector-network-analyzer calibration algorithm of [1] implemented in the MultiCal® software package³ to a multiline TRL calibration algorithm based on the less-well-known iterative approach of [2]. We show that the iterative approach of de-embedding on-wafer scattering-parameter measurements not only outperforms the MultiCal algorithm in the presence of measurement noise, but also accurately estimates the uncertainty of its results.

The multiline TRL algorithm of [1] combines compactness and speed with an effective weighting and averaging strategy based on Gauss-Markov estimates. The algorithm was optimized for on-wafer measurements, and has been incorporated into the convenient and popular MultiCal software package.

The iterative approaches of [2] and [3] offer alternative solutions to the multiline TRL problem based on a nonlinear least-squares solution to the conventional VNA and six-port calibration problems, respectively. While the iterative approaches are slower and less compact than the algorithm of [1], they are designed for optimal performance in the presence of measurement noise.


We later adapted the nonlinear least-squares solution of [2] to the characterization of planar coupled transmission lines in [6-9]. In this case, the least-squares solution was obtained using the orthogonal distance regression algorithm implemented in ODRPACK [10]. The algorithms of [6-9] took advantage of the ability of ODRPACK to determine confidence intervals for the results directly from measurement data.

In this work, we adapt the calibration algorithm of [2] to the orthogonal distance regression algorithm of [10]. As in [2], the new algorithm finds an optimal solution to the multiline TRL on-wafer calibration problem in the presence of random measurement errors. In addition, the new algorithm determines confidence intervals for its results.

In this paper, we demonstrate that this new adaptation of [2] outperforms the multiline TRL calibration algorithm of [1] in the presence of random measurement errors. We also verify the accuracy of its uncertainty estimates.

II. THE CALIBRATION PROBLEM

Figure 1 shows a diagram of the two-tier on-wafer calibration problem that we address with the new calibration algorithm. The matrices \( \mathbf{S}_1 \) and \( \mathbf{S}_2 \) contain the scattering parameters of two microwave ground-signal-ground probe heads to be characterized. The matrix \( \mathbf{S}_C \) contains the scattering parameters of the on-wafer calibration standard contacted by the probes. The elements

³ MultiCal may be obtained at www.boulder.nist.gov/micro
⁴ US government publication, not subject to copyright.

Fig. 1. The on-wafer calibration problem.
of \([S_M]\) are the scattering parameters of the cascade of the two probe heads with scattering-parameter matrices \([S_1]\) and \([S_2]\) and the calibration standard with scattering-parameter matrix \([S_c]\) measured by a network analyzer at the coaxial reference plane indicated in the figure. Here the prime indicates that \([S_M]\) is a measured, rather than a calculated, quantity. The objective of the calibration is to determine the scattering-parameter matrices \([S_1]\) and \([S_2]\) of two probe heads from measurements \([S_M]\) of the probes and on-wafer calibration standards.

In the multiline TRL calibration, the on-wafer standards consist of a short “thru” line, a set of additional on-wafer transmission lines of different lengths, and a symmetric “reflect” [11]. In other on-wafer calibration methods, the lines and/or reflect may be replaced by a variety of previously characterized terminations or other on-wafer calibration standards.

III. THE CALIBRATION ALGORITHM

The orthogonal distance regression algorithm implemented in ODRPACK [10] finds an optimal solution for \(\beta\) of the \(n\) equations

\[
y_i = f_i(x_i + \delta_i, \beta) - \epsilon_i,
\]

where the subscript \(i\) corresponds to the \(i^{th}\) of the \(n\) “observations.” The \(f_i\) are functions relating the measurements \(y_i\) to the unknown vector \(\beta\) and the explanatory variables \(x_i\). The \(\epsilon_i\) and \(\delta_i\) are the errors we wish to minimize in \(y_i\) and \(x_i\).

To solve the calibration problem of Fig. 1, we set elements of the measurement vectors \(y_i\) to the real and imaginary parts of the elements of the measured scattering-parameter matrices \([S_M]\) of the two probes and calibration standard. The vector \(\beta\) contains the unknowns we wish to determine: we assigned elements of \(\beta\) to the real and imaginary parts of the elements of the scattering-parameter matrices \([S_1]\) and \([S_2]\) of the probe heads and, when appropriate, the propagation constant \(\gamma\) of the on-wafer transmission-line standards and the unknown reflection coefficients of any symmetric on-wafer reflect standards.

The vectors \(x_i\) contain sets of “explanatory” variables for each observation. We use them to add previously characterized standards to the calibration, setting elements of the \(x_i\) to the real and imaginary parts of the elements of the scattering-parameter matrix \([S_c]\) of the calibration standard. This strategy not only allows the algorithm to accommodate imperfectly characterized calibration standards, but it allows it to be applied to a broad range of calibration problems, including TRL, open-short-load-thru (OSLT) and line-reflect-match (LRM) calibrations. However, since the TRL calibration does not rely on previously characterized calibration standards, there is no need for the explanatory variables \(x_i\) or their associated errors \(\delta_i\) and weights \(w_\delta\) for that special case.

The optimal solution for \(\beta\) is found by minimizing

\[
\sum_{i=1}^{n} (\epsilon_i^T w_\epsilon \epsilon_i + \delta_i^T w_\delta \delta_i),
\]

subject to the constraints in (1), where the matrices \(w_\epsilon\) and \(w_\delta\) are weights. In our implementation of the on-wafer calibration algorithm, we set \(w_\epsilon\) and \(w_\delta\) equal to estimates of the inverse of the covariance matrices of errors in \(y_i\) and \(x_i\) supplied by the user, which improves the estimate of the unknowns in the vector \(\beta\) obtained with uniform weighting [10].

IV. QUALITATIVE MEASUREMENT COMPARISON

We compared the performance of our new calibration algorithm based on orthogonal distance regression to the algorithm of [1] implemented in MultiCal. Figure 2 compares the magnitude of the transmission coefficient of the first probe head estimated by the two algorithms for one of our typical on-wafer calibrations. The figure shows that the MultiCal estimates are close to the new calibration algorithm’s estimates and, in fact, usually lie well within the standard uncertainty \(s\) as estimated by the new algorithm. We obtained similar results for both the magnitudes and phases of all of the scattering parameters of the probe head.

V. QUANTITATIVE MEASUREMENT COMPARISON

We developed a simulator to examine more closely the performance of the two algorithms. The simulator began with the measured scattering parameters \([S_1]\) and \([S_2]\) of the two probe-heads and propagation constant of the lines,
A. Bias in the Algorithms

Let \( \Delta z \) represent the differences of estimates of \( z \) from the true value of \( z \). The \( t \) statistic \( t = \frac{\Delta z}{s_z} \) is the ratio of the mean of \( \Delta z \), which we write as \( \Delta \bar{z} \), and the standard uncertainty \( s_z \) of \( \Delta z \). Large values of \( t \) indicate significant bias in the estimates of \( z \).

We compiled \( t \) statistics to look for bias in MultiCal and the new algorithm in Table 1, where \( z \) corresponded to estimates of elements of \( [S_k] \) generated by the algorithms. Here \( \Delta z \) corresponds to the difference between (a) the element of \( [S_{k2}] \) listed in the first column of the table determined from the noisy measurements \( [S_{Mk}] \) and (b) the true value of the element in \( [S_k] \). \( \Delta z \) is the mean of the 1000 \( \Delta z \), and \( s_z \) is the standard uncertainty of \( \Delta z \). From the table we conclude that neither algorithm adds a statistically significant bias into its estimates in the presence of Gaussian noise.\(^1\)

Table 1.
The \( t \) statistic for the algorithms. ( \( \sigma_R = \sigma_{R_k} = \sigma_T = \sigma_T = 0.01 \) )

<table>
<thead>
<tr>
<th>( z ) (( \in [S_k] ))</th>
<th>MultiCal</th>
<th>New Algorithm</th>
</tr>
</thead>
<tbody>
<tr>
<td>( \text{Re}(S_{11}) )</td>
<td>0.4</td>
<td>0.7</td>
</tr>
<tr>
<td>( \text{Re}(S_{22}) )</td>
<td>0.02</td>
<td>0.02</td>
</tr>
<tr>
<td>(</td>
<td>S_{21}</td>
<td>)</td>
</tr>
<tr>
<td>Angle ( S_{21} )</td>
<td>0.2</td>
<td>-0.3</td>
</tr>
</tbody>
</table>

B. Variance of the Algorithms

To further explore the performance of the two algorithms, we estimated \( s_{MC} \), the uncertainty in the MultiCal estimates, and \( s_{NEW} \), the uncertainty in the new algorithm’s estimates. We estimated \( s_{MC} \) and \( s_{NEW} \) from the standard deviation of the differences of the 1000 noisy estimates \( [S_{k1}] \) and the true values \( [S_k] \). We tabulated our estimates of \( s_{MC} \) and \( s_{NEW} \), as well as our estimate of \( s_{MC}/s_{NEW} \), in Table 2.

We also list the 95% lower confidence bound \( B_L \) for the ratio based on our estimate of \( s_{MC}/s_{NEW} \) in the last column of Table 2, calculated from the \( F \) distribution with 999 and 999 degrees of freedom using \( B_L = s_{MC} / (s_{NEW} \sqrt{F_{0.05}}) \). There is a 95% certainty that the actual value of the ratio is greater than \( B_L \).

Table 2 shows that the uncertainty \( s_{MC} \) in the MultiCal estimates is consistently greater than the uncertainty in the estimates determined by the new algorithm. Furthermore, the table shows that the new algorithm outperforms MultiCal even when it is not supplied with accurate

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\(^1\) From the \( t \) distribution with 999 degrees of freedom we see that a value of \( \vert t \vert > 1.96 \) is required to conclude with 95% certainty that the estimates have a statistically significant bias.
estimates \( \hat{\sigma}_T \) of the noise in the measurements. Finally, the table indicates that MultiCal has particular difficulty in the presence of noise in transmission-coefficient measurements.

### Table 2.
Uncertainty of the two algorithms. (\( \sigma_r = \sigma_T = 0.01 \))

<table>
<thead>
<tr>
<th>( z ) (in ([S_1]))</th>
<th>( \sigma_T )</th>
<th>( \tilde{\sigma} )</th>
<th>( s_{MC} ) (( \times 10^{-5} ))</th>
<th>( s_{NEW} ) (( \times 10^{-5} ))</th>
<th>( s_{MC}/s_{NEW} )</th>
<th>( B_L )</th>
</tr>
</thead>
<tbody>
<tr>
<td>Re(( S_{11} ))</td>
<td>0.01</td>
<td>0.01</td>
<td>4.96</td>
<td>4.74</td>
<td>1.05</td>
<td>0.997</td>
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<tr>
<td>Re(( S_{21} ))</td>
<td>0.01</td>
<td>0.01</td>
<td>5.50</td>
<td>4.40</td>
<td>1.25</td>
<td>1.19</td>
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<tr>
<td>Re(( S_{11} ))</td>
<td>0.03</td>
<td>0.01</td>
<td>6.36</td>
<td>5.13</td>
<td>1.24</td>
<td>1.18</td>
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<tr>
<td>Re(( S_{21} ))</td>
<td>0.03</td>
<td>0.03</td>
<td>11.8</td>
<td>7.43</td>
<td>1.58</td>
<td>1.50</td>
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#### C. Uncertainty Estimate Generated by the New Algorithm

The new algorithm uses the residual deviations of the measurements from the electrical calibration model to estimate the uncertainties in its own results. Table 3 investigates the accuracy of the standard-uncertainty estimates \( s \) generated by the new algorithm. The table compares the actual standard deviation \( \sigma_{ACTUAL} \) of the quantities in the first column of the table to the mean \( \bar{s} \) of the standard uncertainty estimates \( s \) generated by the new algorithm. The nearly identical values of \( \sigma_{ACTUAL} \) and \( \bar{s} \) indicate that, on average, the new algorithm accurately estimates the uncertainty in its results due to random measurement noise.

The quantity \( u(s)/\bar{s} \) in the last column of Table 3 represents the ratio of the standard deviation \( u(s) \) of the estimates \( s \) to their mean value \( \bar{s} \). The small values of \( u(s)/\bar{s} \) indicate that the new algorithm estimates the uncertainty of its results with reasonable consistency.

### Table 3.
Accuracy of \( s \). (\( \sigma_r = \sigma_T = 0.01 \))

<table>
<thead>
<tr>
<th>( z ) (in ([S_1]))</th>
<th>( \sigma_{ACTUAL} )</th>
<th>( \bar{s} )</th>
<th>( u(s)/\bar{s} )</th>
</tr>
</thead>
<tbody>
<tr>
<td>Re(( S_{11} ))</td>
<td>0.0047</td>
<td>0.0046</td>
<td>0.12</td>
</tr>
<tr>
<td>Re(( S_{22} ))</td>
<td>0.0055</td>
<td>0.0054</td>
<td>0.12</td>
</tr>
<tr>
<td>( S_{21} )</td>
<td>0.0044</td>
<td>0.0044</td>
<td>0.12</td>
</tr>
<tr>
<td>Angle ( S_{21} )</td>
<td>0.324</td>
<td>0.297</td>
<td>0.30</td>
</tr>
</tbody>
</table>

### SOFTWARE

Software implementing this method can be downloaded at http://www.boulder.nist.gov/dylan/.

### REFERENCES


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2. MultiCal estimates only the relative uncertainty in its results as a function of frequency.